



Defining dosing pattern characteristics of successful tapers following methadone maintenance treatment: results from a population-based retrospective cohort study

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ABSTRACT

Aims To identify dose-tapering strategies associated with sustained success following methadone maintenance treatment (MMT). Design Population-based retrospective cohort study. Setting Linked administrative medication dispensation data from British Columbia, Canada, Participants From 25 545 completed MMT episodes, 14 602 of which initiated a taper, 4183 individuals (accounting for 4917 MMT episodes) from 1996 to 2006 met study inclusion criteria. Measurements The primary outcome was sustained successful taper, defined as a daily dose ≤5 mg per day in the final week of the treatment episode and no treatment re-entry, opioid-related hospitalization or mortality within 18 months following episode completion. Findings The overall rate of sustained success was 13% among episodes meeting inclusion criteria (646 of 4917), 4.4% (646 of 14 602) among all episodes initiating a taper and 2.5% (646 of 25 545) among all completed episodes in the data set. The results of our multivariate logistic regression analyses suggested that longer tapers had substantially higher odds of success [12-52 weeks versus <12 weeks: odds ratio (OR): 3.58; 95% confidence interval (CI): 2.76–4.65; >52 weeks versus < 12 weeks: OR: 6.68; 95% CI: 5.13–8.70], regardless of how early in the treatment episode the taper was initiated, and a more gradual, stepped tapering schedule, with dose decreases scheduled in only 25-50% of the weeks of the taper, provided the highest odds of sustained success (versus <25%: OR: 1.61; 95% CI: 1.22–2.14). Conclusions The majority of patients attempting to taper from methadone maintenance treatment will not succeed. Success is enhanced by gradual dose reductions interspersed with periods of stabilization. These results can inform the development of a more refined guideline for future clinical practice.

Keywords Methadone detoxification, methadone maintenance treatment, methadone tapering, opioid dependence, opioid substitution treatment.

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INTRODUCTION

Opioid dependence is a chronic, recurrent disease [1], and long-term patient maintenance on pharmacotherapeutic agents has been recognized as a key component of effective treatment [2]. Methadone maintenance treatment (MMT) is the most common and effective treatment for opioid dependence [3–5], and has been shown to improve health-related quality of life and protect against mortality while clients are engaged in treatment [6,7].

A randomized trial found MMT to be more effective in retaining individuals in treatment, reducing illicit drug use, drug-related human immunodeficiency syndrome (HIV) risk behaviours and illegal activity than long-term detoxification [8]. Systematic reviews of methadone detoxification have indicated that while withdrawal severity can be reduced, the majority of patients relapse to heroin use [9], although success is more likely with additional psychosocial care [10]. In fact, very few patients complete the process of tapering from MMT

[11–15], and relapse is common among those who do reach a state of abstinence [16]. Despite widespread indications for long-term maintenance and largely negative results for detoxification, whether by patient or physician choice, methadone detoxification and dose tapering among maintained clients is observed commonly worldwide [17-19]. Individuals may choose to detoxify due to unrealistic expectations for recovery, pressure from family members or others, the social stigma associated with methadone, programme-related factors, factors associated with financial or travel restrictions and a desire to determine whether clients can cope without treatment [20]. Alternatively, physicians may choose to taper clients for disciplinary reasons or, in some jurisdictions, due to funding constraints or by standardized regulations on clinical practice.

Despite more than 40 years of experience and millions of clients served world-wide, a number of important questions regarding methadone dose-tapering strategies remain. Do the commonly recommended weekly taper guidelines, ranging from 5 to 10% [21–23], achieve optimal outcomes? Should a constant rate of tapering be maintained throughout the tapering process? Rather than considering only the rate of tapering, is the time from treatment episode initiation and the duration of the taper important?

The sparse evidence available provides little insight into effective dosing strategies or other predictors of success. A randomized trial of 63 MMT clients in rapid (10% of initial dose per week) versus gradual (3% of initial dose per week) methadone tapering found gradual tapering to be more effective [24]. Several other shortterm and small-sampled studies identified the length of time in opioid replacement treatment, sustained abstinence from illicit drug use during treatment, vocational/ financial stability, social stability (socialization with nondrug users) and strong motivation to detoxify as factors associated with successful tapering [25-27], although there was wide variation in outcome measurement across studies. A more recent study of 30 maintained clients found no successful methadone tapers, 20 clients stopped their tapers due to drug relapse or medical instability, and 10 remained in treatment over a maximum 41 months of follow-up [12].

Questions regarding clinical guidance for methadone tapering remain unanswered due, in part, to a lack of available longitudinal data and feasibility of long-term experimental studies. Population-level analysis of dosing dynamics during MMT episodes is facilitated in British Columbia (BC), Canada by a centralized medication dispensation database covering the entire province [18,28]. Methadone is recommended to be delivered with a maintenance-oriented philosophy in BC. Treatment is provided by specially licensed physicians in office-based

settings or specialized drug treatment centres, some offering additional psychosocial care. In most cases, medication is dispensed and consumed at community pharmacies under direct supervision. The general ineffectiveness of detoxification treatment is recognized [29], and penalization of patients who relapse into illicit opioid use as well as voluntary withdrawal is discouraged. BC Provincial guidelines for dose tapering recommend a gradual taper at no more than 5% of initial dose per week [21,30].

Our objective was to identify methadone dosetapering strategies associated with sustained success in terms of health and treatment outcomes following maintenance treatment. We conduct a population-level study of methadone clients in British Columbia, Canada to fulfil this objective.

METHODS

Patient population

This study utilized data from the BC PharmaNet database, which records all prescription medication dispensation in the entire province of BC. The cohort was selected from all individuals receiving methadone maintenance treatment (MMT) over an 11-year period: 1 January 1996-31 December 2006. Methadone prescribed for opioid dependence was distinguished from prescriptions for pain indications by separate drug identification numbers in the data set. Data fields available included a de-identified patient ID, date of birth, gender, drug identification number, the date of the prescription ('date'), the length of the prescription ('number of days supplied', or 'days'), medication dosage ('quantity'), de-identified prescriber code and pharmacy code, as well as a geographical identifier, aggregated by local health area. Data on hospitalizations from the BC Ministry of Health Discharge Abstract Database (dates and most responsible diagnoses of hospitalizations) and Vital Statistics (dates and causes of death) were linked to this database throughout the period of study follow-up through Population Data BC [31]. The study was approved by the University of British Columbia/Providence Health Care Behavioural Research Ethics Board.

Measures

Episode definition

MMT episodes were the focal point of our analysis. These were constructed using the 'service date' and 'days supplied' fields; a treatment episode length was calculated as the difference between the last and first days of dispensed medication {episode length = $[(date_{t1} + days_{t1}) - date_{t0}]$ }, where t_0 is the date of

episode initiation, t_1 is last date of service and $days_{t1}$ is the duration of the dispensation at time t_1 within a period of continuous retention in treatment, entailing no interruptions in prescribed doses lasting longer than 30 days. We considered all MMT episodes beginning after 31 January 1996 to ensure consistency in our calculation of episode lengths and eliminate left-censored observations. We focused our analysis on completed treatment episodes, thus also excluding right-censored observations. Efforts were made to correct any misclassification in the data set with regard to prescribed doses, lengths and dates of prescriptions. Fewer than 1% of identified errors could not be corrected.

Taper definition and study selection criteria

Consistent with previous research [18], periods of dose tapering were identified by examining changes in the mean daily dose per week (mean dose: $\sum_{i=1}^{J} \frac{quantity_i}{days_i} / J, J \le 7$ during the course of each episode. The onset of the tapering period was defined as first instance, after the 12-week point, where the weekly dose decreased and remained at or below this lower level for at least 4 weeks $[dose_{t-i} < dose_t \le dose_{t+i},$ i = (1, ..., 4)]. This method was chosen to eliminate the possibility of misclassification of taper initiation as a result of dose titration, vacation catch-up prescriptions, database errors or otherwise. Further, preliminary analysis indicated that many episodes with an observed initiated taper were either reversed or reverted to a stabilized dose. In such cases, in the absence of information of patient or physician intentions (i.e. maintenance dose adjustment or intended taper) we selected completed episodes in which the mean dose was either decreasing over the final 4 weeks of the episode or had decreased to ≤5 mg per day during this period {included if (i) [(mean $dose_{t-3} - mean dose_{T})/mean dose_{T}] < 0$; or (ii) mean dose_t \leq 5 mg, where t = (t-3, t-2, t-1, T) and T is the terminal week of treatment. The durations of tapers were therefore a minimum of 4 weeks.

Sustained successful taper definition

Given that an administrative data source was used, information regarding opioid use beyond the conclusion of treatment episodes was not available, resulting in potential outcome misclassification. We therefore constructed a multi-attribute proxy measure, defined as 'sustained success in tapering' using all available linked administrative data. As successful dose tapering is only the initial requirement towards the ultimate goal of indefinite opioid abstinence [32,33], this measure of sustained success featured four criteria: (i) methadone dose tapered to ≤ 5 mg/day in the final week of treatment, and within

18 months of treatment: (ii) no treatment re-entry; (iii) no drug-related hospitalization; and (iv) no mortality. The 18-month reference period was chosen based on distributions of time to treatment re-entry, to capture the majority of re-entries following successful tapering while maintaining sufficient patient follow-up.

Explanatory variables

As dose control is ultimately at the discretion of the physician, we were specifically interested in how dosing practices influenced the odds of sustained taper success. To this end, individual plots of weekly dose-tapering trajectories were inspected visually to inform construction of relevant covariates capturing dose changes during the tapering period. A set of five covariates were constructed: (i) the taper start week [number of weeks since treatment episode initiation, defined as: mean dose, < (mean dose, 1, . . . , mean dose,); mean dose, \geq (mean dose, , . . . , mean dose,); t \geq 12 weeks]; (ii) the total duration of the taper, from the taper start week to the end of the treatment episode (in weeks); (iii) the median rate of dose change throughout the duration of the taper [taper change rate = median

$$\left[\left(\frac{\textit{mean dose}_t - \textit{mean dose}_{t-1}}{\textit{mean dose}_{t-1}}\right) * 100\right], t = (\text{taper start week},$$

..., episode end week)]; (iv) the median rate of dose change during weeks in which the dose was decreasing throughout the duration of the taper (i.e. taper change rate when $(mean\, dose_t - mean\, dose_{t-1}) < 0$); and (v) the percentage of weeks during which the dose was decreasing throughout the duration of the taper.

Other covariates, added to control for potential confounding, included the maximum mean dose per week (maximum dose) during the treatment episode, treatment adherence [the percentage of missed doses during the

treatment episode:
$$\left(adherence = \frac{\sum_{i=1}^{T} days_i}{episode \ length}\right)$$
, where

T=max (episode length)], age, gender, episode count, calendar year of episode initiation and individuals' Charlson Comorbidity Index (CCI) score [34], based on hospitalizations during the period 6 months prior to treatment episode initiation. Continuous variables were categorized taking into account observed distributions, and with the intention of maximizing clinical interpretability.

Data analysis

Univariate analyses were conducted to identify variables for inclusion in multivariate models. Multivariate logistic regression analysis was used to identify independent factors associated with sustained success. As some individuals had repeated taper attempts, generalized linear

mixed regression models, with logit link and binomial distribution, were also estimated to capture unmeasured individual-specific confounding. Due to the collinearity and conceptual overlap between dose taper dynamics variables (ii), (iii) and (iv–v), three separate multivariate regression models were estimated with covariates (i) and (ii), (i) and (iii) and (i), (iv) and (v), controlling for all other covariates. These sets of variables capture different aspects of the taper, and thus provide complementary evidence on appropriate means of tapering. The goodness-of-fit of these three model formulations were compared using Akaike's and Bayesian information criteria (AIC and BIC). All analysis was executed with SAS version 9.2.

RESULTS

From a total of 32 656 treatment episodes captured in the study, 25 545 were non-censored, from which 14 602 were identified as having initiated a taper at some point during the episode. Among these episodes, 4917 were decreasing or had decreased below 5 mg/day in the final 4 weeks of the episode, thus meeting study inclusion criteria, and ended prior to 1 July 2005, thus allowing 18 months follow-up to verify our definition of sustained success. These episodes were counted among 4183 unique individuals.

From the study sample of 4917 treatment episodes, 1305 (27%) episodes were completed with a daily dose \leq 5 mg/day; however, among episodes reaching this criteria, 659 were followed by at least one of the following events: treatment re-entry within 18 months (458 episodes, 35.1% of all completed tapers), opioid-related hospitalization [319 (24.4%)], mortality [29 (2.2%)]. As such, 1305-659=646 episodes (13% of the study sample) were defined as sustained successful tapers.

Summary statistics stratified by treatment outcome are presented in Table 1. Individuals with treatment episodes resulting in a sustained successful taper were younger, more likely to be male, had lower CCI scores, better treatment adherence, lower maximum mean weekly doses and longer taper durations.

Results of multivariate analyses on the odds of sustained success in the three models tested are presented in Table 2. Patient demographics and treatment adherence variables were included in each model, and their effects were largely consistent across model formulations. Individuals aged 25–34 and 35–49 years had 27–40% lower odds of sustained success in tapering compared to individuals younger than 25 years. Females also had lower odds of sustained success, with the most conservative estimated odds ratio (OR) suggesting 19% lower odds of sustained success compared to males [model 3: OR: 0.81; 95% confidence interval (CI): 0.67–0.99]. Individuals

with higher levels of medical comorbidity also had lower odds of sustained success in tapering; however, this effect was not statistically significant in model 2, when tapering dynamics were represented with the taper duration [0.83 (0.68–1.00)]. Later subsequent treatment episodes had lower odds of sustained success, while episodes initiated in more recent calendar years, controlling for all other measured covariates, had higher odds of sustained success in tapering.

Treatment adherence was included in each of the model formulations, and was associated positively with higher odds of sustained success. Our results suggest that a 1% increase in adherence resulted in a minimum 1.9% increase in the odds of success [model 2: 1.02 (1.00–1.03)].

Results of model 1 suggest that individuals reaching a maximum dose of between 60--100 mg and more than 100 mg had 44 and 60% lower odds of sustained success in tapering compared to those maintained on lower doses. Univariate results were similar to those obtained in the multivariate model. An early taper start week (before week 12) had higher unadjusted odds of sustained success; however, the adjusted effect was not statistically significant. Finally, more aggressive tapers had lower odds of sustained success, as episodes with median percentage decreases greater than 4% were nearly 27% less likely to result in sustained success $[0.73 \ (0.61\text{--}0.88)]$ in comparison to less aggressive tapers.

Model 2 indicated that taper duration was associated strongly with sustained success. Episodes in which the taper lasted 12–52 weeks were 3.58 times more likely to result in sustained success than those lasting <12 weeks [3.58 (2.76–4.65)], while tapers lasting longer than 52 weeks were 6.7 times more likely to result in sustained success [6.68 (5.13–8.70)].

Model 3 provides a refinement of model 1 results, indicating that tapers with 25–50% of weeks where the dose was decreasing had the highest odds of success [1.61 (1.22–2.14)]. In weeks during which the mean dose was decreasing, a median percentage change <5% had lower odds of success [0.53 (0.41–0.71)] than episodes in which the median percentage change was 5–15%, while higher percentage changes in mean dose when the dose was decreasing did not have statistically significantly different odds of sustained success.

Of the three models tested, model 2, with taper dynamics represented by the taper duration covariate, provided the highest AIC and BIC values, indicating greatest model fit and thus highest explanatory power. Each model formulation was tested using a generalized linear mixed regression model; however, given the low number of repeated taper episodes, results using this type of model were identical to the pooled analysis presented.

Table 1 Summary statistics.

	Taper with sustained success ^a	Unsuccessful taper $(n = 4271)$	
	(n = 646)		
CCI score mean (SD)	0.04 (0.43)	0.16 (0.93)	
Episode no., median (IQR)	1 (1, 2)	1 (1, 2)	
Calendar year, median (IQR)	5 (4, 6)	5 (3, 7)	
% Treatment adherence ^b , median (IQR)	96.4 (91.8, 98.5)	94.7 (88.6, 97.5)	
Taper start week, median (IQR)	19 (13, 37)	22 (14, 38)	
Female, no. (%)	184 (28)	1464 (34)	
Age, no. (%): <25	162 (25)	728 (17)	
25–34	217 (34)	1573 (37)	
35–49	231 (36)	1737 (41)	
≥50	36 (6)	233 (5)	
Maximum dose, no. (%) ^c : <60 mg per day	247 (38)	1040 (24)	
60–100 mg per day	287 (44)	2114 (50)	
>100 mg per day	112 (17)	1117 (26)	
Taper length, no. (%): <12 weeks	88 (14)	1829 (43)	
12–52 weeks	238 (37)	1303 (31)	
>52 weeks	320 (50)	1139 (27)	
% change in dose, no. (%): <1% ^d	327 (51)	1909 (45)	
1-4%	81 (13)	548 (13)	
>4%	238 (37)	1814 (42)	
% change in dose when decreasing, no. (%): <5%	65 (10)	695 (16)	
5–10%	283 (44)	1567 (37)	
10–15%	154 (24)	883 (21)	
>15%	144 (22)	1126 (26)	
% weeks dose decreased: <25%	79 (12)	665 (16)	
25–50%	230 (36)	1094 (26)	
50-75%	176 (27)	975 (23)	
75–90%	75 (12)	397 (9)	
>90%	86 (13)	1140 (27)	

^aDefined as an episode in which the daily dose in the final week is ≤5 mg/day, and within 18 months of episode completion, the individual does not re-enter treatment, is not hospitalized for a drug-related cause and does not suffer mortality. ^bPercentage of days received medication during treatment episode. ^cMaximum mean weekly dose through duration of treatment episode. ^dMedian percentage change in mean weekly dose throughout duration of taper. CCI: Charlson Comorbidity Index; IQR: interquartile range; SD: standard deviation.

DISCUSSION

We designed our analysis to inform the development of improved clinical practices for methadone dose tapering. Cumulatively, the results of our multivariate analyses suggest that longer tapers had substantially higher odds of success, regardless of how early in the treatment episode they were initiated, and a more gradual tapering schedule with dose decreases scheduled in only 25-50% of the weeks in treatment provided the highest odds of sustained success in tapering. These results provide a more refined and informative guideline for future practice. However, these results, and methadone dosing guidelines in general, are intended to supplement physician decision-making, rather than replace it. While patient and physician decision-making were not observed in the database, and although such actions were discouraged in clinical guidelines, one might speculate that the more aggressive tapers were disciplinary in

nature. Regardless of the reason for such tapers, it is clear that outcomes are substantially poorer in such cases. Conversely, episodes with more gradual tapering, resulting in higher odds of success, may be indicative of individualized care, with doses titrated and tapered according to individual needs in suppressing withdrawal symptoms: a practice which is widely recommended [2,20–22].

While we have characterized dose-tapering strategies with higher odds of sustained success, the overall rate of success was just 13% among episodes selected for this analysis, 4.4% (646 of 14 602) among all episodes initiating a taper, and 2.5% (646 of 25 545) among all completed episodes in the data set. These poor outcomes are consistent with the findings of previous analyses [9–15], and are observed despite the frequency with which tapers are observed in BC, contrary to clinical guideline recommendations. Further, longer duration of exposure in treatment is associated with improved post-treatment

Table 2 Results of multivariate regression analysis of sustained success in dose tapering.

	Unadjusted odds ratios	Adjusted odds ratios	Adjusted odds ratios	Adjusted odds ratios
		Model 1	Model 2	Model 3
Covariate	Odds ratio (95% CI)	Odds ratio (95% CI)	Odds ratio (95% CI)	Odds ratio (95% CI)
Age: <25	Ref	Ref	Ref	Ref
25–34	0.62 (0.50-0.77)	0.71 (0.56-0.89)	0.78 (0.61-0.98)	0.72 (0.57-0.91)
35–49	0.60 (0.48-0.74)	0.69 (0.55-0.87)	0.76 (0.60-0.96)	0.70 (0.55-0.89)
≥50	0.69 (0.47-1.03)	0.71 (0.47-1.06)	0.79 (0.52-1.21)	0.74 (0.49-1.12)
Gender: male	Ref	Ref	Ref	Ref
Female	0.76 (0.64-0.92)	0.79 (0.65-0.95)	0.75 (0.61-0.91)	0.81 (0.67-0.99)
CCI score	0.75 (0.62-0.91)	0.82 (0.67-0.99)	0.84 (0.69-1.03)	0.82 (0.67-1.00)
Episode no.	0.71 (0.64-0.78)	0.72 (0.65-0.80)	0.73 (0.65-0.81)	0.73 (0.66-0.81)
Calendar year	1.03 (0.99-1.07)	1.08 (1.04-1.13)	1.09 (1.05-1.15)	1.07 (1.02-1.12)
Treatment adherence ^a	1.04 (1.03-1.05)	1.04 (1.02-1.05)	1.02 (1.00-1.03)	1.03 (1.03-1.05)
Maximum dose ^b : <60 mg per day	Ref	Ref		Ref
60–100 mg per day	0.57 (0.48-0.69)	0.56 (0.46-0.67)	0.46 (0.37-0.56)	0.58 (0.48-0.71)
>100 mg per day	0.42 (0.33-0.54)	0.39 (0.31-0.51)	0.31 (0.23-0.40)	0.43 (0.33-0.55)
Taper start week	1.28 (1.03-1.59)	1.01 (1.00-1.01)	1.00 (1.00-1.00)	1.00 (1.00-1.01)
Taper length: <12 weeks	Ref		Ref	
12-52 weeks	3.80 (2.94-4.90)		3.58 (2.76-4.65)	
>52 weeks	5.84 (4.56-7.48)		6.68 (5.13-8.70)	
% change in dose: <1% ^c	Ref	Ref		
1–4%	0.86 (0.67-1.12)	0.77 (0.59-1.01)		
>4%	0.77 (0.64-0.92)	0.73 (0.61-0.88)		
% change in dose when decreasing: <5%	0.52 (0.40-0.69)			0.53 (0.40-0.71)
5-15%	Ref			Ref
15–35%	0.75 (0.60-0.93)			1.03 (0.82-1.31)
35-50%	0.70 (0.43-1.16)			1.42 (0.83-2.43)
>50%	0.42 (0.18-0.97)			1.12 (0.47-2.67)
% weeks dose decreased: <25%	Ref			Ref
25-50%	1.77 (1.35-2.33)			1.61 (1.22-2.14)
50-75%	1.52 (1.14-2.02)			1.29 (0.96-1.73)
75–90%	1.59 (1.13-2.23)			1.32 (0.92-1.87)
>90%	0.63 (0.46-0.87)			0.56 (0.40-0.79)
AIC	_	3631.82	3485.36	3569.91
BIC	_	3714.14	3548.12	3684.65

Ref: reference group; 95% CI: 95% confidence interval; AIC: Akaike's information criterion; BIC: Bayes information criterion; CCI: Charlson Comorbidity Index. "Percentage of days received medication during treatment episode. "Maximum mean weekly dose through duration of treatment episode. "Median percentage change in mean weekly dose through duration of taper. Blank cells indicate that the variable was not included in the specified multivariate regression model.

outcomes such as reduced opioid use, reduced criminal activity, improved social productivity and reduced risk of mortality [29,35–38]. While our focus in this paper has been on the technical aspects of therapeutic tapering, there a number of reasons why methadone maintenance clients may wish to achieve a drug-free state in the process of recovery from drug use [32,33]. However, the high risks of drug relapse and adverse health outcomes need to be communicated to clients choosing to do so. In light of these risks, an informed consent process should be considered by regulatory bodies before initiating a tapering schedule.

The negative relationship between the maximum, or maintenance, dose, as well as results of later subsequent taper treatment episodes and the probability of success need to be interpreted carefully. A recent Cochrane Review of dosing in MMT programmes concluded that clients receiving maintenance doses of 60–100 mg per day were retained in treatment longer than those on lower doses [39], while a previous analysis of this data set confirmed this result and found maintenance doses >100 mg/day were associated with longer retention in treatment [28]. Clinical guidelines in BC instruct physicians to use discretion, and work with patients to provide individualized care in identifying a suitable dose which eliminates withdrawal symptoms and drug craving. That said, physician decision-making was not observed, and methadone dose is often determined by the policies of

individual practices. In cases where doses were titrated to individuals' needs, those requiring higher daily doses could be considered to have higher addiction severity requiring a higher maintenance dose and, perhaps not surprisingly, had lower odds of successfully tapering. Although the relationship between the maintenance dose and a prior drug use severity is not observed and therefore not testable in our data set, the notion that higher a priori drug use severity decreases the likelihood of successful transition to abstinence is supported in the available literature [40,41].

Strengths and weaknesses

There are notable limitations that are important to consider when interpreting results of this study. First, the outcome we have constructed based on the data available to us is a proxy of sustained abstinence, and cannot capture those who revert to illicit opioid use without returning to treatment or experiencing the adverse health outcomes included in the outcome. What is being assessed is therefore more indicative of medical safety and patient stability following treatment rather than drug relapse. Further study is needed to confirm these results prospectively and potentially over longer time-frames. Secondly, due to the nature of the data set, physician decision-making was not observed, which required us to classify stages of the methadone treatment episodes, particularly the point of initiation of a taper, according to observed changes in doses. This needs to be considered carefully in further research and practical application. Thirdly, any non-experimental study may be subject to residual and/or unmeasured confounding [42]—in our case, we could not observe whether clients were primarily users of heroin or prescribed opioids, addiction severity. treatment setting (office-based versus drug treatment centre) or availability of additional psychosocial care and other potential confounders. In a separate analysis, we used the random-effects frailty term from a previous analysis on the time to discontinuation of repeated methadone treatment episodes as an additional covariate in each model to determine the potential effects of individual-level unmeasured confounding. Results were similar to those presented, with only minor differences in the magnitude of effect sizes. However, we are careful not to interpret the tested relationships as causal, as these unobserved factors may confound or modify the effects of dose changes on sustained success. Finally, the results in the current study represent outcomes of practices in BC, Canada where treatment is heavily subsidized, a maintenance-oriented philosophy persists and officebased treatment is available. In other settings, barriers due to financial constraints or availability may impact clinical practice and compromise treatment outcome.

The advantages of using this data set, however, were substantial: availability of a centralized drug dispensation database allowed for a population-level study of methadone dose tapering among maintained clients unparalleled in size. The large sample size and level of detail in methadone dosing allowed us to classify and compare dose-tapering strategies, providing new insights into clinical practice. We are hopeful that this study will spur further research on tapering strategies for clients in opioid substitution treatment programmes wishing to achieve a state of opioid abstinence.

Declarations of interest

None.

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